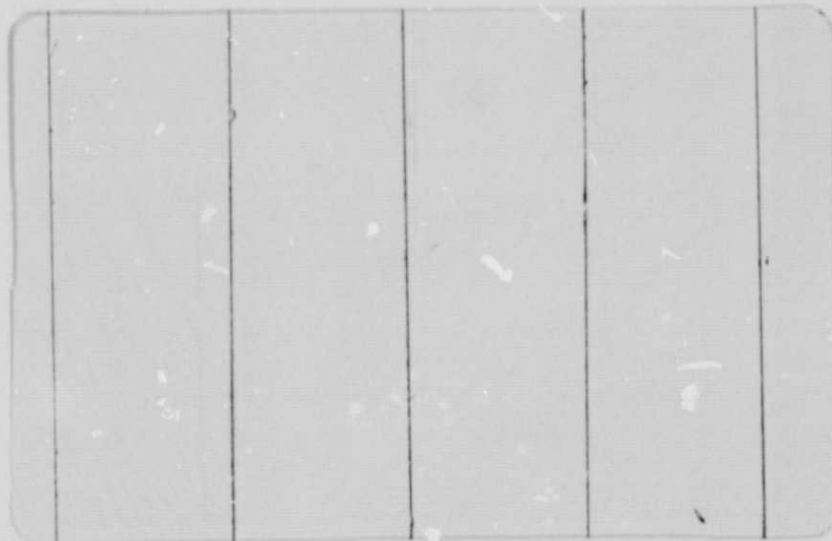


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The Recursive Maximum Likelihood
Proportion Estimator—User's Guide
and Test Results

by

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ABSTRACT:

In this report, we describe our implementation of the recursive maximum likelihood proportion estimator proposed by D. Kazakos in "Recursive Estimation of Prior Probabilities Using the Mixture Approach," (Rice University, ICSA Technical Report #275-025-019). A user's guide to the programs as they currently exist on the IBM 360/o7 at LARS, Purdue is included, and test results on LANDSAT data are described. On Hill County data, the algorithm yields results comparable to the standard maximum likelihood proportion estimator.

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I. Introduction:

In this report, we describe our implementation of the recursive maximum likelihood proportion estimator proposed by D. Kazakos in [1]. Numerical results obtained with this algorithm using LANDSAT data are described, and a user's guide for the programs as they currently exist on the IBM 360/67 at LARS (terminal available at NASA-JSC) is included.

Section II contains a description of the algorithm as implemented. Section III serves as a user's guide to the programs available. In section IV, we describe the numerical results we have obtained with this algorithm. An appendix contains listings of the programs.

II. The Algorithm:

Given a set of n -dimensional measurement vectors $\{x\}$ from M normally distributed multivariate pattern classes H^j , $j=1, 2, \dots, M$ the $M-1$ dimensional recursive maximum likelihood proportion estimate (RMLPE)⁽¹⁾ p^i at the i^{th} data vector is given by

$$p^i = p^{i-1} + \frac{1}{i} L \left[g(p^{i-1}, x_i) \right]^{-1} \\ \left(f_1(x_i) - f_M(x_i), f_2(x_i) - f_M(x_i), \right. \\ \left. \dots, f_{M-1}(x_i) - f_M(x_i) \right) \quad (1)$$

where $f_j(x)$ is the density function for the j^{th} class;

$$f_j(x) = (2\pi)^{-n/2} |K_j|^{-\frac{1}{2}} \exp \left[(x - u_j)^T K_j^{-1} (x - u_j) \right] \quad (2)$$

where u_j and K_j are the mean and covariance matrix, respectively, for the j^{th} class; $g(p^{i-1}, x_i)$ is the mixture distribution estimate, i.e.,

$$g(p^{i-1}, x_i) = f_M(x_i) + \sum_{\ell=1}^{m-1} p_{\ell}^{i-1} (f_{\ell}(x_i) - f_M(x_i)) \quad (3)$$

and L is a suitably chosen constant in this approximation. The proportion estimate for the M^{th} class is denoted by p_m^i and given by

$$p_m^i = 1 - \sum_{\ell=1}^{m-1} p_{\ell}^i \quad (4)$$

In our implementation of this algorithm, we have made several modifications to improve its performance. These include (1) clipping the value of the update (i.e. second) term in eq. (1); (2) renormalizing the p^i at each step so that all $p_j^i \geq 0$ and $\sum_{\ell=1}^m p_{\ell}^i = 1$; and (3) introducing an additional damping term in the update term of eq.(1). The final form of the algorithm is

$$p^i = \text{NORM} \left\{ \epsilon, p^{i-1} + \frac{1}{i+n_0} \text{LMT} \left[T, L g(p^{i-1}, x_i) \cdot \left(f_1(x_i) - f_M(x_i), f_2(x_i) - f_M(x_i), \dots, f_{M-1}(x_i) - f_M(x_i) \right) \right] \right\}$$

where $\text{LMT}(a, b)$ is the clipping function defined by

$$\text{LMT}(a, b) = \tilde{b}$$

with
$$\tilde{b}_i = \text{sign}(b_i) \min(a, |b_i|)$$

NORM is the renormalizing function defined by

$$\text{NORM}(\epsilon, y) = \text{the first } M-1 \text{ elements of } \tilde{y}$$

where

$$\tilde{y}_m = 1 - \sum_{i=1}^{m-1} y_i$$

$$\tilde{y}_i = y_i$$

If

$$\min(\tilde{y}_i) \geq \epsilon > 0 \text{ then finish else}$$

$$\tilde{y}_i \leftarrow \tilde{y}_i - \min(\tilde{y}_i) + \epsilon \quad i = 1, 2, \dots, M$$

$$\tilde{y}_i \leftarrow \tilde{y}_i / \sum_{\ell=1}^m \tilde{y}_\ell ;$$

and n_0 is a positive constant used to damp out early oscillations of the estimate.

Two other algorithms used in conjunction with this one are (1) an algorithm to calculate an approximation to L and (2) an algorithm to scramble all of the data (the RMLPE uses the stochastic approximation, so the data needs to appear in a random order). The first algorithm calculates the following approximation to L

$$L = \left(u \cdot \min |K_j|^{\frac{1}{2}} \right)^{-1}$$

where u is the minimum eigenvalue of H with $H = \{h_{ks}\}$ and

$$h_{ks} = (2\pi)^{n/2} \int_{E^n} (f_k(x) - f_m(x)) (f_s(x) - f_m(x)) dx$$

$$k, s = 1, 2, \dots, m$$

The scrambling algorithm employs a procedure described on page 125 of [2].

III. Program Description and Users Guide

Three programs have been written to implement this algorithm: the proportion estimation program, a program to calculate an approximation to L and a program to scramble data prior to estimating proportions. These programs are described below and listings are provided in the appendix.

Proportion Estimation Program:

This program runs on the IBM 360/67 at LARS. Parameters are read from cards describing characteristics of the data and the statistics, the processing to be performed, and the desired outputs. The same data may be processed with several sets of statistics. The data is assumed to be sixteen channel data (from which any subset of channels may be used) residing on file 11 with one logical record per data vector. The data may be labelled or unlabelled. If labelled, the program will calculate the true proportions and print out the means of the estimates along with the associated variances and mean squared error; if unlabelled, the true proportions are read from cards and the same quantities are then computed. Due to the use of the stochastic approximation in

this algorithm, the data vectors must be scrambled before being put on file 11. (Program super AM may be used for this purpose.)

The correspondence between the notation used in the preview section and the variables in the program are:

<u>Above</u>	<u>Program</u>
p_i^j	Q(J, *)
n_o	ISTRT
ϵ	EPS
T	TLM
L	L
G	G
$f_i(x)$	F (a function subprogram)
\sum_i	SG
u_i	MU

The programs are set up to handle up to 16 channel data from up to 15 classes with as many as 10 different blocking factors. They can treat an unlimited number of data points. The data enters the program in "lines" which contain ≤ 1500 points.

[N.B. The total number of points need not be an integral multiple of the points per line,
 e.g. if there are 5100 total points, we may
 use NP (the number of points/line = 200
 and NL (the number of lines) ≥ 26]

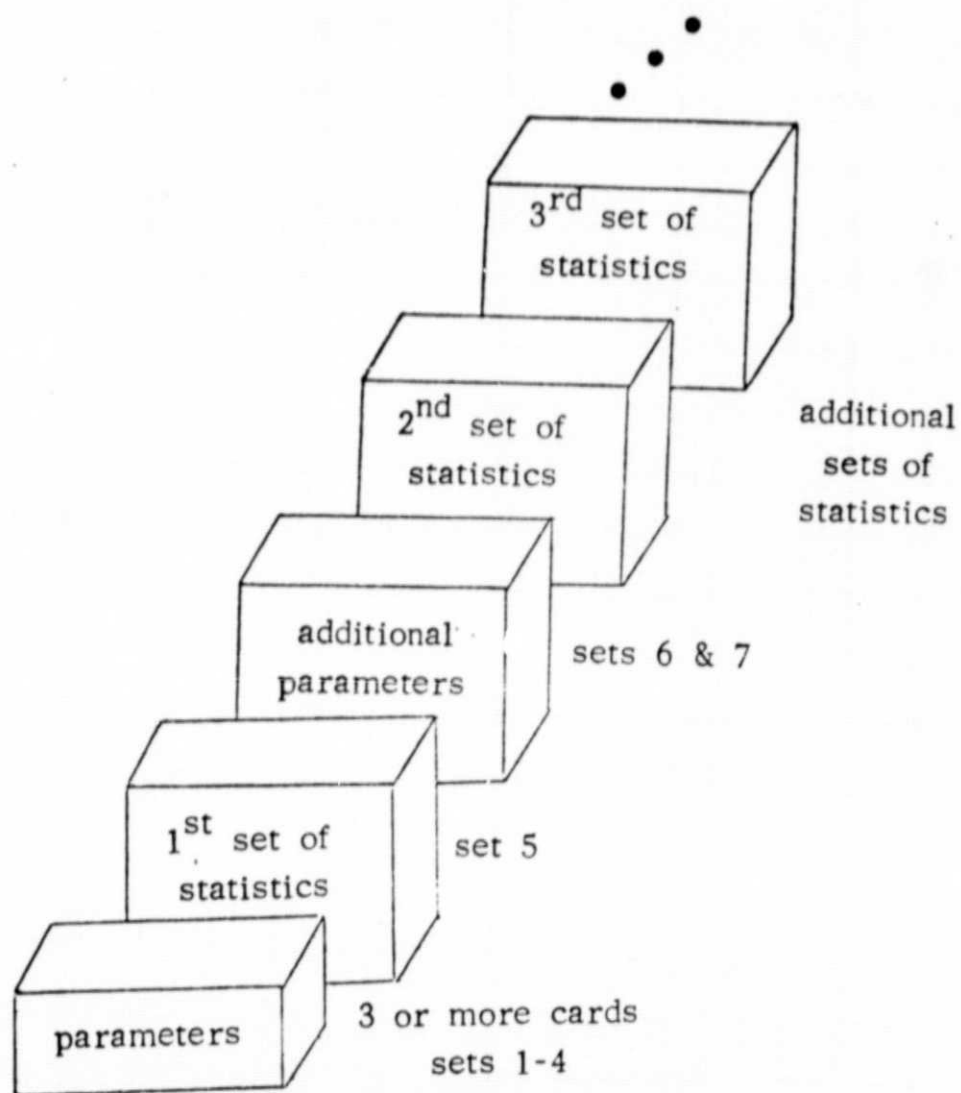


Figure 1

Data Deck Setup for the LANDSAT Version

Figure 1 shows the set-up of the data deck necessary to execute the program. The input parameters and their formats are described below:

- 1) HEDNG - Title to be printed on the output (20A4)
- 2) M, MXITER, NK, ISTRT, INQ, ØUTPT,
L, TLM, EPS, (K(I) , I=1, NK)

(4 I 2, 2 L 1, 3 G 10.8, 10 X, 10 I 3)

- M - number of classes used
- MXITER - number of sets of statistics to use
- NK - number of blocking factors to use (set = 1)
- ISTRT - starting value of n_0 in eq. (1)
(default = 99)
- INQ - = F if the initial guess for the proportion estimates (Q's) are to be set = $\frac{1}{M}$ (then card set (3) are not used)
= T if the Q's are to be read in (card set (3) is required)
- ØUTPT - = T if updated Q's are to be printed after each line of data. Otherwise set = F
- L - the L value to be used in eq. (1)
- TLM - the maximum permissible absolute value for the update quantity for the Q's
(i.e.

$$L = \sum_{\ell=1}^K \frac{f_j(x_j) - f_m(x_s)}{G(p_{i-1}, x_s)}$$

$$s = K * (i-1) + \ell$$

EPS - minimum allowable value for a Q during the estimation procedure (10^{-2} seems to be a good choice)

$K(I), I=1, NK$ - the blocking factors to be used (set $K(1)=1$)

3) CSET
(16 L 1) - $CSET_i = \begin{cases} T & \text{if } i^{\text{th}} \text{ channel is to be used} \\ F & \text{otherwise} \end{cases}$

Optional 4) $((Q(I, J), I=1, M-1), J=1, NK)$
(16 G 5.3)

The initial guess for the Q 's. Used only if INQ on card 2 is = T

5) $CL(I), (MU(J, I), J=1, 16)$
(26 X, A 1/(5 X, 5 E 15.8)
(SG (J, I), J=1, 136)
(5 X, 5 E 15.8) } for $I=1, 2, \dots, M$

These cards contain the statistics for the M classes. CL is the class ID, MU, the mean vector, and SG is the covariance matrix stored in symmetric storage mode (i. e. upper triangular part stored by columns). Note that there are 33 cards required for each class. Additional sets of statistics follow card set (7).

6) NP, NL, OUTPP, OUTPX, TRUEP
(2 I 5, 3 L 1)

NP - number of points to use per "line" (≤ 1500)

NL - maximum number of "lines" of data

\emptyset UTPP - = T the current true proportions are
 printed after each line (used only if
 TRUEP = F)
 = F do not print these proportions
 \emptyset UTPX - = T print the data vectors
 = F do not print the data vectors
 TRUEP - = T if the true proportions are to be read
 in (card set (7) is then required)
 = F the class ID is associated with each
 data vector and the program will calculate
 the true proportions (card set (7) not used).

Optional 7) (CLS(J), GT(J), J=1, M)
 (8(A 2, G 8.6))

CLS(J) - the class ID for the J^{th} class

GT(J) - the true proportions for the J^{th} class

The data vectors should be on file 11 with 1 data vector per logical record in the format

CL, (X(J), J=1, 16)

(8 X, A 1, 6 X, 16 F 4.0, 1 X)

where CL is the class ID (used only if TRUEP on card set 6 = F) and X(J) contains the 16 dimensional data value for a pixel.

The subroutines used in this program are briefly described below:

INSTAT - reads and prints statistics

SUBSET - for LANDSAT data (LD) version, this selects appropriate subsets of the statistics.

- F - computes the value of the density function at X
- TPØSE - for the pseudo-random (PR) version, transposes the data matrix in situ
- GEDATA - obtains or generates a line of data in the required format and order. Also computes the true proportions.
- MCHLSK - computes the modified Cholesky decomposition of a covariance matrix stored in symmetric storage mode.

Program to Calculate L:

This program calculates the following approximation to L

$$= \left(h \cdot \min_j \left| \sum_j \right|^{\frac{1}{2}} \right)^{-1}$$

where $h = \min \left(\text{eval} (H) \right)$

$$\begin{aligned} & \& H_{ks} = (2\pi)^{n/2} \int_{E^n} \left(f_k(x) - f_m(x) \right) \cdot \\ & \quad \cdot \left(f_s(x) - f_m(x) \right) dx \\ & \quad k, s, = 1, 2, \dots, m \end{aligned}$$

All notation is as before.

Input parameters to the program are

CSET, M, N

(16L1, 2X, 12, 2X, 12)

[N.B. Our (limited) experience with the proportion estimation algorithm indicates that a value of ~ 3 for L appears optimal despite what this program computes.]

Scrambling Program:

This algorithm scrambles the order of records in a data set and creates a new data set. Two storage arrays are used: one containing the integers $1, 2, \dots, N$ where N is the total number of records and the other containing space for one data record. A temporary direct access data set, which is the same size as the original data set, is used. The algorithm is described below:

- 1) Set $a_i = i$ for $i = 1, 2, \dots, N$
- 2) Scramble the elements of the vector a .
(see e.g. ref. [2]).
- 3) For $i = 1, 2, \dots, N$
 - a) Read i^{th} record of original data set and store it in vector d .
 - b) Write d in a_i^{th} record in temporary data set.
- 4) For $i = 1, 2, \dots, N$
 - a) Read i^{th} record of temporary data set and store it in vector d .
 - b) Write d on i^{th} record of new data set.
- 5) Finished.

Note that step 4 may not be necessary if one can use the data from the temporary direct access data set.

IV. Numerical Results:

A variety of numerical experiments were conducted with this program to determine its characteristic. Both pseudo-random and LANDSAT data were used.

The most significant effect of this algorithm is due to the scrambling (i.e. the order in which the data is input). If the data is not scrambled (i.e. blocks of points from single classes appear to the program) unreliable estimates will be produced. Our experience with LANDSAT data indicates that the entire data set, whose proportions are to be estimated, needs to have the individual pixels scrambled. Various scramblings will produce different estimates with a theoretical variance of L/N where N is the total number of pixels.

Another effect that we noticed was that the variance of the estimate for the M^{th} class was always larger than for other classes. This asymmetry, we feel, is due to the fact that the algorithm estimates proportions for the first $M-1$ classes, and the estimate for the M^{th} class is then computed as $1 - \sum_{i=1}^{M-1} p^i$. By reordering the classes and then again estimating proportions, it was determined that the variance of the M^{th} class would decrease from $\sim 10\%$ to $\sim 30\%$, so the effect may not be too harmful. However, the user should be aware of this and assure that the estimate for the M^{th} class is of the least interest.

Detailed tests of this algorithm were run on some Hill County LANDSAT data in order to compare results with those obtained by Coberly and Odell [3] with five other proportion estimation algorithms. Table 1 shows the results obtained from the recursive maximum likelihood estimator (RMLE) for 2600 pixels of the labelled data as compared to the other five estimators. Note that the RMLE and MLE have almost equal variances and mean squared errors.

Table II shows the results obtained from the RMLE for 8400 pixels of the unlabelled data. Here again the variances and mean squared error are approximately the same as those of the MLE.

V. Conclusions:

Our experience with this algorithm indicates several important factors need be taken account of in using this algorithm: (1) all of the data needs to be scrambled point by point, (2) the class of least importance should be used as the last class, and (3) a value of ~ 3 for the parameter L appears close to optimal.

Our tests indicate that the recursive maximum likelihood estimator (RMLE) produces results of comparable variance and accuracy as the standard maximum likelihood estimator (MLE) of ref. [3]. The amount of computation involved for the RMLE is equivalent to the first iteration of the MLE plus the scrambling of the data. Also, no additional storage is required by this algorithm to store the density functions for each data point.

Further tests of this algorithm with other LANDSAT data will be necessary to determine the effectiveness of this algorithm in the general situation.

Table 1
Summary of Experiment I
(Labeled Data, 2600 Pixels)

	CLASS	ODELL	MLE	RMLE	MIX	MCM	GT
MEAN	WH	.297041	.300439	.307807	.294352	.274749	.371900
	FA	.274428	.296764	.277315	.312696	.235257	.286200
	BA	.174306	.176688	.188004	.183300	.197059	.115400
	GR	.085066	.085825	.058506	.074200	.075168	.079200
	ST	.168155	.140178	.168364	.135448	.217765	.147300
VAR	WH	.000174	.000083	.000094	.000201	.000459	
	FA	.001385	.000408	.000347	.002216	.003027	
	BA	.000190	.000084	.000125	.000123	.000288	
	GR	.000051	.000060	.000050	.000165	.000463	
	ST	.001829	.000383	.000533	.002376	.004145	
TOTAL VAR		.003528	.001017	.001149	.005020	.008383	
MSE		.013322	.010086	.010778	.016572	.032066	

Table 2
Summary of Experiment II
(Total Data Set, 8400 Pixels)

	CLASS	ODELL	MLE	RMLE	MIX	MCM	GT
MEAN	WH	.252733	.262044	.272700	.267325	.226467	.084333
	FA	.207300	.183526	.185133	.197484	.353019	.023415
	BA	.167633	.154997	.151200	.142416	.218460	.325809
	GR	.186466	.187198	.189166	.158599	.158955	.184282
	ST	.186066	.212231	.201700	.234173	.043098	.382159
VAR	WH	.000095	.000253	.000310	.000307	.000232	.000949
	FA	.000269	.007430	.000472	.000550	.002872	.002114
	BA	.000063	.000299	.000206	.000220	.000135	.000610
	GR	.000261	.000363	.000529	.000609	.001042	.000900
	ST	.000141	.009569	.000832	.000903	.002975	.003402
TOTAL VAR		.000829	.017924	.002350	.002590	.007257	.007974
MSE		.008622	.057982	.010273	.008176	.055378	.180347

REFERENCES

- [1] D. Kazakos, "Recursive Estimation of Prior Probabilities Using the Mixture Approach," ICSA Technical Report #275-025-019, Rice University, Houston, Texas, September, 1974.
- [2] D. Knuth, Seminumerical Algorithms, the Art of Computer Programming, Vol. 2, Addison-Wesley, Reading, Mass., 1969, p.125.
- [3] W.A. Coberly and P.L. Odell, "An Empirical Comparison of Five Proportion Estimators," from the annual report of the University of Texas at Dallas for NASA contract NAS 9-13512, January, 1975.

APPENDIX

FILE. . . DVR FORTRAN P1

```

REAL*4 SG(136,30),MU(16,30),O(30,10),L,X(16,1500),DET(30),
1 T(30,10)/300*0./
REAL*4 DUM(2700)
REAL*4 A(30),G(10)
REAL*4 OS(30,10),OB(30,10)/300*0./,OV(30,10)/300*0./
REAL*4 GT(15)
REAL*4 MSE(30)
REAL*4 HEDNG(20)
REAL*8 S,SS
INTEGER*4 K(10),IP1(10)/10*0/
INTEGER*4 CHAN(16)
INTEGER*2 CL(15)
LOGICAL*1 INQ,IENP,IND,OUTPT,END
LOGICAL*1 CSET(16)
LOGICAL*1 FRST/.TRUE./
COMMON /PASS/ SG,MU,M,N
COMMON /RSFT/ NSET,GT
COMMON /GEPTS/ NL
COMMON /ORDR/ CL
NAHFLIST /IDAT/ M, NK,ISTRT,IND,OUTPT,L,TLM,FPS,XXM,MXITER,CSET
MXCHN=16
NXPTS=1500
MXCLS=30
C
C M - NUMBER OF CLASSES USED (.LE.30)
C N - NUMBER OF CHANNELS USED (.LE.16)
C MXITER - NUMBER OF TIMES TO REDO THE RUN WITH DIFFERENT DATA
C NK - NUMBER OF K'S TO BE USED
C ISTRT - INITIAL VALUE OF J IN O(R+1)=O(R)-1/(J+R)*L*F. (DEF=10)
C INQ - LOGICAL VARIABLE INDICATING WHETHER TO READ INITIAL
C GUESS FOR THE O'S OR NOT
C OUTPT - LOGICAL VAR =T IF ESTIMATE OF O IS TO BE PRINTED
C AFTER EACH LINE OF DATA
C L - THE L VALUE USED BY THE ALGORITHM
C TLM - LIMITING VALUE OF L*ABS(EJ-EM)/GI (DEF=INFINITY)
C FPS - LOWER LIMIT ALLOWED FOR THE O'S
C XXM - UPPER LIMIT ALLOWED FOR THE O'S
C CSET - ARRAY INDICATING WHICH OF THE 16 CHANNELS ARE TO BE USED
C K - THE BLOCKING FACTORS TO BE USED (.LE. 10 OF THEM, EACH LE NPD)
C O - THE ESTIMATES OF THE PRIORS
C SG - COVARIANCE MATRICES STORED IN SYM STORAGE MODE
C MU - MEAN VECTORS
C X - THE DATA VECTORS FOR 1 'LINE' OF DATA
C
C PI2=2.*3.14159265
C END=.FALSE.
C
C READ PARAMETERS
C
C READ (5,1) HEDNG
C 1 FORMAT (20A4)
C WRITE (6,4) HEDNG
C 4 FORMAT (//1X,20A4,///)
C READ (5,2) M,MXITER,NK,ISTRT,IND,OUTPT,L,TLM,FPS,XXM,(K(I),I=1,NK)
C 2 FORMAT (4I2,2L1,4G10.8,10I3)
C READ (5,3) CSET
C 3 FORMAT (16L1)
C IF (ISTRT.EQ.0) ISTRT=10
C IF (TLM.LT.1.E-2) TLM=1.E70
C M1=M-1
C ITER=0
C IF (INQ) GO TO 5
C GO TO 10
C
C THEN READ IN INITAAL GUESS FOR PRIORS FOR EACH BLOCKING USED.
C
C 5 READ (5,7) ((O(I,J),I=1,M1),J=1,NK)
C 7 FORMAT (16G5.3)
C GO TO 20
C
C ELSE SET INITIAL GUESS FOR PRIORS ALL EQUAL
C
C 10 Y=1./M
C DO 15 I=1,M1
C DO 15 J=1,NK
C 15 O(I,J)=Y
C
C *****
C

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C
20 CONTINUE
  DO 26 I=1,M)
    DO 26 J=1,NK
26  OS(I,J)=O(I,J)
    WRITE (6,10AT)
    TLM=TLM/L
    WRITE (6,21) (K(I),I=1,NK)
21  FORMAT (' K=',10I5)
    DO 23 I=1,NK
23  WRITE (6,22) (O(J,I),J=1,M1)
22  FORMAT (' INITIAL O=',8G16.8)
  GET STATS FOR THE CLASSES
C
28 CONTINUE
  CALL INSTAT (CSET, CHAN)
  IF (IFR.FO.O) PI2=PI2**(N/2.)
  DO 34 II=1,NK
34  IP1(II)=O
C
  DO CHOLFSKY DECOMP OF THE COVARIANCES
C
  DO 25 I=1,M
    CALL HCHLSK (SG(1,I),N,X,DET(I))
25  DET(I)=1.DO/(SORT(DET(I))*PI2)
C*****
C
  FETCH ONE LINE OF DATA
C*****
60 CALL CADATA(X,NP, CHAN,N,8100)
C
  IF (.NOT.ERST) GO TO 32
  ERST=.FALSE.
C
  LOOP OVER ALL DATA POINTS TO UPDATE ESTIMATE OF PRIORS
C
32 CONTINUE
  IENP=.FALSE.
  DO 30 I=1,NP
    IF (I.EQ.NP) IENP=.TRUE.
C
  F YIELDS THE VALUE OF THE DENSITY FUNCTION FOR THE CLASS
C
  FM=F(X(1,I),N,SG(1,M),MU(1,M),DET(M))
C
  THE MIXTURE DISTRIBUTION IS SCORED IN G
C
  DO 35 II=1,NK
35  G(II)=FM
    DO 40 J=1,M1
      FJ=F(X(1,I),N,SG(1,J),MU(1,J),DET(J))
      AJ=FJ-FM
      G(II)=G(II)+O(J,II)*AJ
40  CONTINUE
C
  LOOP TO UPDATE PRIORS FOR EACH BLOCKING FACTOR
C
  DO 50 II=1,NK
    IND=.FALSE.
    KI=K(II)
    MIK=MOD(I,KI)
    GI=G(II)
    IF (MIK.EQ.O.OR.IENP) GO TO 52
    GO TO 53
C
  THEN PREPARE TO UPDATE II-TH PRIORS
C
52 IK=MINO(MIK,KI)
  IF (IK.EQ.O) IK=KI
  IND=.TRUE.
  IP1(II)=IP1(II)+IK
53 CONTINUE
C
  COMPUTE UPDATED SUMS

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S=0.
DO 55 J=1,M1
XX=A(J)/G1
T(J,II)=T(J,II)+SIGN(AMIN1(ABS(XX),TLM),XX)
IF (.NOT.IND) GO TO 56
GO TO 55
C
C UPDATE THE PRIORS AND RESET
C
56 XX=O(J,II)+L*T(J,II)/(IP1(II) +ISTRT)
T(J,II)=0.
O(J,II)=XX
S=S+XX
55 CONTINUE
IF (.NOT.IND) GO TO 50
C
C RENORMALIZE THE UPDATED ESTIMATES OF THE PRIORS
C
O(M,II)=1./DO-S
S=O(1,II)
DO 54 J=2,M
IF (O(J,II).LT.S) S=O(J,II)
54 CONTINUE
IF (S.GT.EPS) GO TO 64
SS=0.
DO 58 J=1,M
O(J,II)=O(J,II)-S+EPS
58 SS=SS+O(J,II)
DO 62 J=1,M1
62 O(J,II)=O(J,II)/SS
64 CONTINUE
50 CONTINUE
30 CONTINUE
C*****
IF (.NOT.OUTPUT) GO TO 60
C
C PRINT OUT NEW ESTIMATE OF PRIORS
C
72 DO 70 II=1,NK
S=0.
DO 75 J=1,M1
S=S+O(J,II)
O(M,II)=1./DO-S
70 WRITE (6,76) NSET,K(II),(CL(J),O(J,II),J=1,M)
76 FORMAT (/' UPDATED ESTIMATE OF THE PRIORS FOR LINE ',I5,' WITH BL
1CKING FACTOR K=',I3/4(' CLASS=',A2,' O=',G15.8,3X))
IF (.NOT.END) GO TO 60
C
C UPDATE MEANS AND VARIANCES OF THE ESTIMATES FOR THIS ITERATION.
C
ITER=ITER+1
DO 116 II=1,NK
DO 115 J=1,M
XX=O(J,II)
OB(J,II)=OB(J,II)+XX
OV(J,II)=OV(J,II)+XX*XX
115 O(J,II)=OS(J,II)
116 CONTINUE
NSET=0
END=.FALSE.
IF (ITER.LT.MXITER) GO TO 28
C
C FINISHED WITH ALL DATA. PRINT OUT ESTIMATES & STOP
C
DO 113 II=1,NK
S1=0.
DO 119 J=1,M
MSE(J)=(OV(J,II)-2.*GT(J)*OB(J,II))/MXITER+GT(J)*GT(J)
S1=S1+MSE(J)
XX=OB(J,II)/ITER
OB(J,II)=XX
XX=OV(J,II)/ITER-XX*XX
119 OV(J,II)=SORT(XX)
WRITE (6,131) K(II),(CL(J),OB(J,II),OV(J,II),MSE(J),J=1,M)
131 FORMAT (' K=',I5,' MEANS AND SD'S OF THE ESTIMATES & THE MSE'/
1 (A3,G16.8,' +- ',G16.8,G18.8))
S1=S1/M
WRITE (6,121) S1

```

(iii)

FILE. . . DVR FORTRAN P1

```
118 CONTINUE
121 FORMAT (' *** MEAN MSE',G16.8)
STOP
C
C FINISHED WITH ALL DATA FOR THIS ITERATION
C
100 END=.TRUE.
GO TO 72
END
```

DVR02350
DVR02360
DVR02370
DVR02380
DVR02390
DVR02400
DVR02410
DVR02420
DVR02430

```
FUNCTION F(X,N,L,MU,DET)
C
C COMPUTE THE VALUE OF THE DENSITY FUNCTION AT X
C
REAL*4 X(1),L(1),MU(1),Y(16)
REAL*8 TF,S
C
C SOLVE L Y=X-MU WHERE L IS THE CHOLESKY DECOMP OF COVAR MATRIX.
C DIAG ELEMENTS OF L ARE STORED AS RECIPROCALLS.
C
S=X(1)-MU(1)
Y(1)=S
TF=S*S*L(1)
IF (N.EQ.1) GO TO 15
K=1
C
C LOOP TO COMPUTE Y(I'S)
C
DO 10 I=2,N
S=X(I)-MU(I)
JJ=I-1
DO 20 J=1,JJ
K=K+1
20 S=S-L(K)*Y(J)
K=K+1
Y(I)=S
TF=TF+S*S*L(K)
10 CONTINUE
15 CONTINUE
IF (TF.LT.325.) GO TO 17
F=0.
RETURN
17 F=EXP(SNGL(-TF/2.))*DET
RETURN
END
```

DVR02440
DVR02450
DVR02460
DVR02470
DVR02480
DVR02490
DVR02500
DVR02510
DVR02520
DVR02530
DVR02540
DVR02550
DVR02560
DVR02570
DVR02580
DVR02590
DVR02600
DVR02610
DVR02620
DVR02630
DVR02640
DVR02650
DVR02660
DVR02670
DVR02680
DVR02690
DVR02700
DVR02710
DVR02720
DVR02730
DVR02740
DVR02750
DVR02760
DVR02770
DVR02780

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SUBROUTINE INSTAT (CSET, CHAN)
REAL*4 SG(136,30),MU(16,30)
INTEGER*4 CHAN(1)
INTEGER*2 CL(15),NC(15),NP(15)
LOGICAL*1 CSET(16)
COMMON /PASS/ SG,MU,M,N
COMMON /OROR/ CL
DO 5 I=1,M
  READ (5,1) CL(I),(MU(J,I),J=1,16)
1  FORMAT (26X,A1/(5X,5F15.8))
5  READ (5,4) (SG(J,I),J=1,136)
4  FORMAT (5X,5F15.8)
  CALL SUBSET(CSET,CHAN)
  DO 10 I=1,M
    WRITE (6,2) CL(I),NC(I),NP(I),(MU(J,I),J=1,N)
2  FORMAT (/ / ' CLASS',A2,I5,' NO. OF PTS=',I5/' MEAN',10F11.4/
1  1X,6F11.4)
    WRITE (6,3)
3  FORMAT (' COVARIANCE')
    J1=1
    J2=0
    DO 20 J=1,N
      J2=J2+J
      J1=J1+J-1
20  WRITE (6,21) (SG(L,I),L=J1,J2)
21  FORMAT (/ (1X,13F10.4))
10  CONTINUE
  RETURN
END

```

DVR02790
DVR02800
DVR02810
DVR02820
DVR02830
DVR02840
DVR02850
DVR02860
DVR02870
DVR02880
DVR02890
DVR02900
DVR02910
DVR02920
DVR02930
DVR02940
DVR02950
DVR02960
DVR02970
DVR02980
DVR02990
DVR03000
DVR03010
DVR03020
DVR03030
DVR03040
DVR03050
DVR03060
DVR03070

```

SUBROUTINE SUBSET (CSET, CHAN)
LOGICAL*1 CSET(1)
REAL*4 SG(136,30),MU(16,30)
INTEGER*4 CHAN(1)
COMMON /PASS/ SG,MU,M,K
ISB(I,J)=(I*(I-1))/2+J
IDAG(I)=(I*(I+1))/2

C
C
C  FIND CHANNELS DESIRED
C
K=0
DO 10 I=1,16
  IF (.NOT.CSET(I)) GO TO 10
  K=K+1
  CHAN(K)=I
10  CONTINUE

C
C
C  SELECT APPROPRIATE SUBSETS OF SG & MU
C
JJ=0
DO 20 I=1,K
  STORE DIAGONAL ELEMENTS
  JJ=JJ+1
  DO 25 L=1,M
    SG(JJ,L)=SG(IDAG(CHAN(I)),L)
25  MU(I,L)=MU(CHAN(I),L)
    IF (I.EQ.K) RETURN

C
C
C  MOVE ALL ELEMENTS OF NEXT ROW EXCDPT THE DIAGONAL ONE
C
L1=CHAN(I+1)
DO 30 J=1,I
  L2=CHAN(J)
  JJ=JJ+1
  DO 30 L=1,M
30  SG(JJ,L)=SG(ISB(L1,L2),L)
20  CONTINUE
STOP
END

```

DVR03080
DVR03090
DVR03100
DVR03110
DVR03120
DVR03130
DVR03140
DVR03150
DVR03160
DVR03170
DVR03180
DVR03190
DVR03200
DVR03210
DVR03220
DVR03230
DVR03240
DVR03250
DVR03260
DVR03270
DVR03280
DVR03290
DVR03300
DVR03310
DVR03320
DVR03330
DVR03340
DVR03350
DVR03360
DVR03370
DVR03380
DVR03390
DVR03400
DVR03410
DVR03420
DVR03430
DVR03440
DVR03450
DVR03460
DVR03470
DVR03480

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	SUBROUTINE GEDATA (X,NP,CHAN,KK,*)	DVR03490
	REAL*4 SG(136,30),MU(16,30)	DVR03500
	REAL*4 X(16,1500),GT(15)	DVR03510
	INTEGER*4 CHAN(1)	DVR03520
	INTEGER*2 CL,CLS(15),PTS(15)/15*0/,ITER/0/,SCLS(15)	DVR03530
	INTEGER*2 LPTS(15)/15*0/,IPTS(15)/15*0/	DVR03540
	LOGICAL*1 FRST/.TRUE./	DVR03550
	LOGICAL*1 OUTPP,OUTPX	DVR03560
	LOGICAL*1 TRUEP	DVR03570
	FOR COBERLY'S DATA	DVR03580
	COMMON /PASS/ SG,MU,M	DVR03590
	COMMON /RSFT/ LINE,GT	DVR03600
	COMMON /OROR/ SCLS	DVR03610
	COMMON /GEPTS/ NL	DVR03620
	IF (.NOT.FRST) GO TO 10	DVR03630
	FRST=.FALSE.	DVR03640
	LINE=0	DVR03650
	K=0	DVR03660
	NP = NUMBER OF POINTS PER 'LINE' (.LE.1500)	DVR03670
	NL = NUMBER OF LINES	DVR03680
	OUTPP=T - PRINT RUNNING TRUE PROPORTIONS	DVR03690
	OUTPX=T - PRINT DATA VECTORS	DVR03700
	TRUEP=T - READ IN TRUE PROPORTIONS	DVR03710
	READ (5,1) NP,NL,OUTPP,OUTPX,TRUEP	DVR03720
1	FORMAT (2I5,3L1)	DVR03730
	NPS=NP	DVR03740
	WRITE (6,2) NP,NL,OUTPP,OUTPX,TRUEP	DVR03750
2	FORMAT (' NP=',I5,' NL=',I5,' OUTPP=',L8,' OUTPX=',L8,' TRUEP=',	DVR03760
1	L8)	DVR03770
	IF (TRUEP) READ (5,3) (CLS(J),GT(J),J=1,M)	DVR03780
3	FORMAT (8(A2,G8.6))	DVR03790
10	LINE=LINE+1	DVR03800
	NP=NPS	DVR03810
	IF (LINE.LE.NL) GO TO 20	DVR03820
	FINISHED WITH THIS PASS OF THE DATA	DVR03830
	REWIND 11	DVR03840
	ITER=ITER+1	DVR03850
	IF (ITER.GT.1) RETURN 1	DVR03860
	COMPUTE TRUE PROPORTIONS & REARRANGE CLASSES TO HHOSE IN STATS	DVR03870
	IF (TRUEP) K=M	DVR03880
	JJ=0	DVR03890
	DO 50 I=1,K	DVR03900
	IF (TRUEP) PTS(I)=0	DVR03910
	DO 55 J=1,K	DVR03920
	IF (CLS(J).EQ.SCLS(I)) GO TO 52	DVR03930
55	CONTINUE	DVR03940
	WRITE (6,53) SCLS(I)	DVR03950
53	FORMAT (' CLASS NOT FOUND',A3)	DVR03960
	GO TO 50	DVR03970
52	IF (I.EQ.J) GO TO 50	DVR03980
	L=CLS(I)	DVR03990
	CLS(I)=CLS(J)	DVR04000
	CLS(J)=L	DVR04010
	IF (TRUEP) GO TO 50	DVR04020
	L=PTS(I)	DVR04030
	PTS(I)=PTS(J)	DVR04040
	PTS(J)=L	DVR04050
50	CONTINUE	DVR04060
	IF (TRUEP) GO TO 57	DVR04070
	DO 54 I=1,K	DVR04080
	JJ=JJ+PTS(I)	DVR04090
54	CONTINUE	DVR04100
	PRINT OUT PROPORTIONS	DVR04110
	57 CONTINUE	DVR04120
	XJ=JJ	DVR04130
	WRITE (6,51)	DVR04140
		DVR04150
		DVR04160
		DVR04170
		DVR04180
		DVR04190
		DVR04200
		DVR04210
		DVR04220
		DVR04230
		DVR04240

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15 CONTINUE
   KK(J1) = TF
12 CONTINUE
   DET = DET * TF
   IF (L.GT.NV) GO TO 10
   IRD = J1 + 1

```

```

25 CONTINUE
16 KK (IRD+J)=T1/TF
20 CONTINUE
   JF1=.FALSE.
10 CONTINUE
   J1=0

```

```

      DO 30 J=1,NV
      J1=J1+J
30    KK(J1)= 1./KK(J1)
      RETURN
      END

```

(viii)

COMPILER OPTIONS -

NAME= MAIN,OPT=CC,LINECNT=60,SIZE=CCOCK,
SOURCE,LCCDIC,NCLIST,NOLCK,LCCAD,MAP,NODEDIT,NOID,XREF

THIS PROGRAM SCRAMBLES NP RECORDS (.LE.10000) EACH OF LENGTH 19
WCROS CONTAINED ON FILE 11 AND PUTS THE RESULTS ON FILE 12

ISN 0C02' INTEGER*2 INT(10000),DAT(19)
ISN 0C03 DEFINE FILE B(10000,10,0,1J)
ISN 0C04 ISEED=1314159793
ISN 0C05 READ (5,11) NP
ISN 0C06 11 FORMAT (15)

GENERATE THE INTEGERS 1,2,...,NP AND STORE IN ARRAY INT

ISN 0C07 DC 40 I=1,NP
ISN 0C08 20 INT(I)=I

SCRAMBLE ARRAY INT

ISN 0CC9 J=NP

GGUFB GENERATES A RANDOM NUMBER FROM U(0,1)

ISN 0C10 25 R=GGUFB(ISEED)
ISN 0C11 IY=J*R+1
ISN 0C12 IT=INT(J)
ISN 0C13 INT(I)=INT(IX)
ISN 0C14 INT(IX)=IT
ISN 0C15 J=J-1
ISN 0C16 IF (J.GT.1) GO TO 23
ISN 0C17 WRITE (6,7)
ISN 0C18 7 FORMAT (: SHFLD*)
ISN 0C19

LOOP TO PUT 1-TH RECORD IN INT(I)-TH POSITION ON FILE 3 (A TEMP.
DIRECT ACCESS FILE)

ISN 0C20 DC 30 I=1,NP
ISN 0C21 L=INT(I)
ISN 0C22 FIND(B,L)
ISN 0C23 READ (11,1) DAT
ISN 0C24 1 FORMAT (2I3,2A1,10I3)
ISN 0C25 30 WRITE (8,L) DAT
ISN 0C26 WRITE (6,8)
ISN 0C27 8 FORMAT (: ON 8*)

COPY FILE 8 TO FILE 12

ISN 0C28 DC 40 I=1,NP
ISN 0C29 I1=I+1
ISN 0C30 READ (8,1) DAT
ISN 0C31 FIND(B,I1)
ISN 0C32 IF (1/50*50.EQ.1) WRITE (6,3) I
ISN 0C33 3 FORMAT (18)
ISN 0C34 40 WRITE (12,1) DAT
ISN 0C35 STOP
ISN 0C36 END
ISN 0C37

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